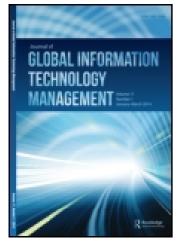
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Exploring Break-points and Interaction Effects Among Predictors of the International Digital Divide

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Exploring Break-points and Interaction Effects Among Predictors of the International Digital Divide

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ABSTRACT

The deepening of the digital divide between countries has prompted international organizations and governments to work together toward reducing the problem over the next 15 years. However, such efforts will likely succeed only if they are based on a firm grasp of the divide's underlying causes. In this paper we report the results of a comprehensive analysis of the determinants of the international digital divide. Our results confirm many findings of past research, but also extend existing knowledge in important ways. By employing Multivariate Adaptive Regression Splines (MARS), we discover non-linearities and interaction effects among the predictors. We then articulate significant policy implications based upon these findings.

KEYWORDS

International Digital Divide, Global IT, ICT, MARS

INTRODUCTION

At the UN Millennium Summit, among several other goals, the authorities pledged to reduce the international digital divide by 2020. This pledge came at a time when the digital divide across countries, following several years of widening, had reached a new extreme. The digital divide, both at the international and the national levels, has three dimensions: one dimension that reflects the basic access to and use of technology, another that reflects the factors that affect the use of technology, and a third that reflects the advanced application of technology. As articulated by the non-profit organization *bridges.org*, these dimensions are reflected in alternate measurements of the digital divide, including a) those quantifying the number of people with *access* to the technology – which in this case is measured as the use of the

Internet, phones, or computers; b) those quantifying the *quality* of the technology – such as the Internet bandwidth, the human capital available in the economy to allow for technical training and use of Information and Communication Technologies (ICT); c) those reflecting the *intensity of usage* of the technology – such as the extent of e-commerce and industrial information technology available in the economy; and finally d) those reflecting the *market* for technology – such as the pricing of ICTs. In this paper, we undertake a comprehensive analysis of the determinants of the international digital divide (IDD). We define IDD broadly as the gap between countries in their level of utilization of information and communication technologies (ICTs).

Studying the digital divide is important for two complementary reasons. First and foremost, there is a sizable literature recognizing the link between Internet connectivity and economic and social wellbeing of a population. Second, understanding the causes and the evolution of disparities in Internet connectivity will facilitate the design and implementation of alternative policy tools that can improve the economic and social wellbeing of the laggard countries and assist them in closing the gap. This paper serves the purpose of identifying the underlying reasons for the existing IDD using improved econometric techniques and data that are more recent and complete.

Distribution of the Digital Divide

The global map of the digital divide has evolved since the introduction of the Internet in North America and Europe starting in 1980s. According to the G8 Digital Opportunity Task Force (DOT Force) the percentage of Internet users in the population of developing countries increased from 0.1 percent to 1.6 percent from 1995 to 2001 (bridges.org, 2006). However, the increase was much more marked for developed countries, rising from 4 percent to 28 percent over the same time span. Figure 1 illustrates the spatial expression of this divide, with OECD countries (Europe, North America, and a handful of Asia-Pacific countries) leading the world in Internet connectivity, followed by several Latin American, Eastern European, and Middle Eastern countries as well as China and South Africa. Extreme laggard regions include Sub-Saharan Africa and other landlocked and impoverished states in South America, Central Asia, and Southeast Asia.

The determinants of IDD have been explored in prior literature, both from a global point of view as well as a regional perspective. Regardless of the scale of analysis, however, common explanations of the IDD can be divided into three categories: economic factors (level and equality of wealth and income), cultural factors (religion and language), and factors related to the telecommunications infrastructure (ownership, infrastructure, pricing). The regional studies furthermore take into account region-specific factors. For example, Sub-Saharan Africa exhibits unique conditions (such as the extent of HIV/AIDS) with respect to IDD, to warrant the generation of a specialized regional literature (see, for example, Mutula, 2004). The Eurasian transition economies also exhibit regional commonalities – such as a particularly wide range of cultural traits – that have justified a regional research approach (Dimitrova and Beilock, 2005).

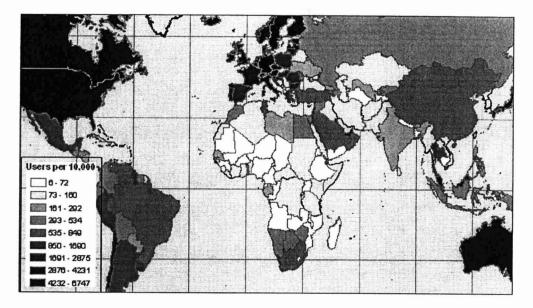


Figure 1. Distribution of the International Digital Divide, 2003. Data Source: World Bank Development Indicators CD-ROM (2005).

Generally speaking, the digital divide between developing and developed countries is widening. Only a few countries, such as China and South Korea, have been able to keep pace with ICT developments but even these have been unable to close the gap because more developed countries continue to move forward at a faster pace (Campbell, 2001).

In a study of ICT production and diffusion in Asia, Wong (2002) examined the extent to which the presence of ICT goods manufacturing has helped Asian countries generate sufficient "digital dividends" to attain higher rates of ICT diffusion. Using regression analysis, the study shows that the link between ICT production and diffusion is rather weak and that Asian countries as a group exhibit higher disparity in ICT adoption than non-Asian countries, after controlling for the level of economic development. Similarly, in a study of ICT industry development in China, Meng (2002) provides empirical evidence that in spite of China's astonishing pace of progress in its ICT industry, a clear digital divide exist among its economic regions.

While within-country digital divide is not visible in Figure 1 and does not lie within the scope of this project, other research highlights intra-national digital divides within countries including Chile (Mendoza and Toledo 1997) and Germany (Kubicek 2004), to name just two examples. Importantly, the findings include the observation that within-country digital divide along socio-economic demographic lines (gender, race, ethnicity, age, income, education) is shrinking in developed countries. For example, recent evidence from the United States suggests that African-Americans, Hispanic Americans and Asian-Americans are fast closing the digital divide by gaining access

to and increasing their use of the Internet (Marriott, 2006). Recent international data collected and compiled by the World Internet Project report a similar trend in several developed countries. However, experts quoted in Marriott (2006) caution that a new dimension of digital divide may be opening up in terms of the type and quality of access to the Internet that is far more challenging than the access alone. Finally, recent evidence from the Oxford Internet Survey (Dutton, 2005) as well as the World Internet Project suggests that the diffusion of the Internet is likely to be reaching a plateau in developed countries. In other words, unlike other technologies, such as TV, the Internet diffusion will in all likelihood reach a point of saturation.

In summary, while Internet access and use continue to increase exponentially on a global basis, the extent of Internet diffusion and utilization varies widely across regions, countries, social classes, and ethnic groups. Research evidence presented above points to a widening gap between developed and developing countries. While the digital divide presents a threat, it also offers tremendous opportunities for developing nations to narrow the social and economic inequalities by facilitating and encouraging a faster diffusion of ICT. In fact, some Southeast Asian countries, such as Thailand and Malaysia, have created new agencies to speed the implementation of ICT and to address digital divide issues (Tipton, 2002). But, it is important to keep in mind that the impact of the Internet diffusion on the economic and social wellbeing of nation is not uniform as technologies interact with a wide range of cultural and social factors to produce certain consequences (Light, 2001).

Determinants of the Digital Divide

Determinants of *inequality in access to ICT* are known as first order effects, while those impacting *inequality in the ability to use ICT* are known as second order effects (Hargittai 2002, Dewan and Riggins 2005). Among the first quantitative attempts to explore and explain the international digital divide is Eszter Hargittai's (1999) paper that reports findings on eighteen members of the Organization for Economic Cooperation and Development. Evaluating a wide range of (mainly first order) variables at the macro level, Hargittai concludes that the most important determinants of international disparities in ICT include economic wealth and telecommunications infrastructure, findings that would be echoed in nearly every subsequent study. irrespective of data scope and timeframe. Specifically, Hargittai (1999) finds that income alone explains 38 percent of the variation of Internet penetration among OECD countries. The "global" analysis of Internet users per 10,000 individuals ("IUR", Internet User Rate) in 105 countries undertaken by Beilock and Dimitrova (2003) also yields the primacy of per capita income as a major determinant of the IDD. In an important regionally-focused follow-up paper on the determinants of the digital divide, Dimitrova and Beilock (2005) reiterate the importance of infrastructure and income in influencing the IDD. Specifically, they find that among the transition economies 81% of the variation in the IUR is explained by income and infrastructure.

In a comprehensive analysis of the determinants of the global digital divide, Chinn and Fairlie (2004) find parallel results to Hargittai (1999), supporting that economic

wellbeing and infrastructure have significant influence on the digital divide. Chinn and Fairlie (2004) indeed find that infrastructure, measured as the telephone density and the regulatory quality, alongside the economic wellbeing of the countries, measured as income per capita, are the leading determinants of computer use. At the same time, the authors discover that Internet penetration is chiefly driven by electricity consumption, youth dependency, income per capita, and regulatory quality, and qualify other studies that identify the importance of telephone density and Internet access pricing. Chinn and Fairlie (2004) underline the non-significance in their models of the years of education and the negative effect of trade openness on Internet penetration. Our methods shed light on these findings. Finally, the authors provide evidence regarding the robustness of their results across different regions of the world.

In examining other first order effects (access to ICT), Hargittai (1999) tests whether the national distribution of income, alongside average income level of the economy, plays a role in affecting the IDD. Her results suggest that it is the level of income rather than the distribution of it --measured by the Gini coefficient-- that matters for the existence and evolution of the IDD (named after the Italian statistician Corrado Gini, the Gini coefficient measures the extent of inequality in a distribution). Testing second order effects (ability to use ICT), Hargittai (2002) later augments her macroeconomic findings at the individual-level, revealing the importance of online skills.

In a more recent paper, Bagchi (2005) examines the factors contributing to the digital divide between members of the Organization for Economic Cooperation and Development (OECD) and Economic Commission for Latin America and the Caribbean (ECLAC). The author defines the digital divide as the division between those who have access to information and communications technology and those who do not. Multiple yardsticks are used in the study to measure this divide, including telephone, cell phone, PC, and Internet usage per 1,000 individuals, from which the author extracts a factor named "IT Index". This index is then correlated with independent variables in line with past research, and linear regression models are developed, confirming earlier results. However, one limitation of this study is the small sample size, due to missing values for a number of indicators.

Chen and Wellman (2004)Nua.com use data from (available from http://www.clickz.com/ at the time of this writing) and the International Telecommunication Union (ITU) to describe changes over time in the percent of population on-line in eight countries. Notwithstanding the shortcomings of their dataset, they find that in just one year the divide between "first movers" (USA) and relative "latecomers" (UK, Korea, and Japan) decreases most rapidly, pointing out that some diffusion continues to take place in laggard states such as China (where the population online increased from 2.5% to 4.8%). Although they do not provide evidence through any modeling or any econometric tests, the authors speculate that the driving forces in the digital divide across countries include cost differentials, English language competency, relevant content, and technological support. However, Hargittai (1999) undertakes a formal testing of the significance of several of these variables, specifically the use of English language and cost differentials (measured as

the telecom policy and pricing), and finds that these two variables are not statistically significant in explaining the variation in IDD.

The IDD literature is not limited to macroeconomic factors that enable ICT. Other extensive portions take an individual level or organizational level approach to analyzing characteristics that affect access to and use of ICT, as reviewed extensively in an organizational framework set forth by Dewan and Riggins (2005). In her pioneering work, Norris (2001) discusses such factors and the international digital divide at the national and regional level. She employs OLS regression analyses to reveal that in Europe, the same social and demographic factors that explain online participation also help predict access to new (computer related) and old (cable and television) media technologies. These include individuals' characteristics such as age, gender, income, and education. Despite the micro level evidence of the important role played by education, in a formal testing of the hypothesis using macro level data. Hargittai (1999) finds that education does not statistically significantly affect the IDD. Norris (2001) furthermore identifies the influence of the institutional structure of the economy on the digital divide by studying how the digital divide interacts with a variety of political participation modes. Using national case studies and data through 2000, she examines connections between social and political as well as economic development variables and Internet use (measured by percent on-line). Interestingly the variable on the level of democratization is not significant once variables on economic and social development and regional dummy variables are controlled for.

In parallel, Beilock and Dimitrova (2003), focus on the former socialist countries of Eastern Europe, former Soviet Republics, and Mongolia, using a multiple regression analysis. Along with the measures of income and infrastructure, they assess the impact of religion and civil liberties on the number of Internet users per 10,000 individuals. In their 2003 study, the authors find that the R-square of their model increases from 77% to over 90% with the addition of a variable measuring Roman Catholicism or Orthodoxy, which yields a positive effect. In their 2005 study, Dimitrova and Beilock confirm the positive effect of Roman Catholicism or Orthodoxy, as well as openness (measured by the level of civil liberties in the country). They also find that the effect of income is insignificant when controlling for infrastructure (income and infrastructure are heavily correlated, at least for the countries involved in this study).

By underscoring the importance of income, even among wealthy states, the above papers support the call for further investigation in particular with reference to the unique characteristics of each country. In response, we set out to investigate a list of variables inspired by this literature, while recognizing *a priori* that the relationship between the IDD and each of its drivers could very likely be non-linear.

Finally, in their comprehensive and timely literature review on the International Digital Divide, Dewan and Riggins (2005) offer a conceptual framework for organizing past studies on the variables impacting ICT access at the individual and organizational level. The authors systematically list all major past studies on ICT adoption, efforts to bridge the digital divide, and patterns of ICT usage and summarize

their key findings, ranging from individuals and households to companies and demographic groups. The result is an overarching conceptualization of the literature on the levels of individual, organizational, and global analysis, the latter being instrumental in guiding the present study. The authors further segment findings into first order (access to ICT) and second order (ability to use ICT) digital divide. Especially with reference to the second-order divide, the authors call for policy recommendations on the productive use of ICT as well as skills that are complementary to this use.

As detailed above in our literature discussion, existing studies on the IDD do not examine such nonlinearities in the relationship between the IDD and its predictors and do not explore the possibility of interactions among predictors, which would imply that the effect of one predictor on the number of Internet users depends on another predictor. By employing the methodology of Multiple Adaptive Regression Splines (MARS), this study explicitly accounts for and empirically identifies the nonlinearities and interactions in the relationship. This contributes to a better understanding of the drivers of the IDD and thus an opportunity to improve policies for closing it.

METHODS AND DATA

Our dependent variable is the number of Internet users per 10,000 individuals (years 2001-2003), as reported by the World Bank. We chose this variable in large part because we possess a relatively complete and reliable dataset for 160 countries, but also because similar yardsticks are widely used in the literature. The measure closely mimics Hargittai's (1999) variable on connected computers per 10,000 and Dimitrova and Beilock's (2005) employment of Internet users per 1000 (or "IURs" – Internet User Rates). Other dimensions of the digital divide can be considered; while the number of Internet users per 10,000 individuals is a measure of basic access to the technology, other dimensions could address the quality of that access, such as bandwidth or human capital available in the economy to support the technology. In order not to overload the paper (and because we lack data, for example, on bandwidth), we will focus here on basic access measured by the number of Internet users per 10,000 individuals. The choice of independent variables included in the analysis was discussed in our literature review. These variables can be categorized into three groups:

- 1. economic, social and political;
- 2. ICT pricing-related variables;
- 3. ICT policy and infrastructure-related variables.

The variables, their time coverage and the sources of the data are listed in Table 1.

The group of economic variables includes variables already discussed in the literature, such as the income level and the income GINI coefficient, as well as variables such as the trade openness (a measure of the importance of trade relative to the size of the economy), the demographic structure, and the urbanization of the country. The role of

international trade in allowing for technology transfers, which could be interpreted as allowing for reducing the international discrepancy in the use of ICT across nations, has been discussed extensively in the literature. Saggi (2002) and Connolly (2003), among others, suggest the possible role of imports in allowing for the economy to benefit from foreign technologies. Our analysis takes a similar perspective and tests the possibility that increased trade among countries could contribute to reducing the IDD by transferring technologies across countries. The importance of age, gender, education and culture were also discussed in the more micro-level studies discussed above. The macro analysis in this paper controls for such effects by including measures of education level, the demographic structure and urbanization level of the population, in addition to the extent of ethno-linguistic division within each society.

Variable	Description	Year(s)	Source
Internet	Number of Internet users per 10,000	2001-03	ICU
Computers	Number of computers per 100 people	2001-03	ICU
School	Average years of schooling	2001-03	World Bank
Gini	Average Gini index for reported years	1989-93	Various ⁱ
Costcall	Cost of three-minute local call (\$PPP)	2001-03	ICU
Electric	Electricity consumption kwh/capita	2001-03	World Bank
GNI	GNI per cap. in international ppp	2001-03	World Bank
Maintel	Number of main telephone lines per 100 p.	2001-03	World Bank
p1564	Percentage of population age 15-64	2001-03	World Bank
p65plus	Percentage of population 65 and older	2001-03	World Bank
Trade	Trade in goods as a percentage of GDP	2001-03	World Bank
urban	Urban population as percent of total	2001-03	World Bank
ethnic	Index of ethnic fragmentation	2001-03	World Bank
risk	Composite Risk Rating Index	2001-03	PRS Group

Table 1. Explanation of Variables

In addition to such measures that reflect economic and social characteristics, the analysis will include a measure of the "risks" in the economy. The composite risk measure (CRRI – Composite Risk Rating Index) not only captures the cyclical economic risks but also the institutional and political soundness of the economies. The higher its value, the less "risky" a country is.

The final group of variables includes those related to the ICT pricing, policy, and infrastructure. We follow the literature by including not only measures of the pricing

policy, but also the available infrastructure measured both as the number of main telephone lines and the electricity consumption per capita.

Our data are from three years: 2001, 2002, and 2003, and represent the most recent virtually complete data set for the 160 countries under investigation. We do not estimate a time lag for our model because this would greatly reduce our sample size. Moreover, a lag would have little impact because most of our independent variables vary only slightly from year to year. Because the years 2002 and 2003 are nearly identical with respect to the relationship between the our dependent and independent variables, we combine the two years as one in our analysis, indicated by the suffix "02" in our tables (this implies one indicator variable for 2002 and 2003 used as a predictor, represented by one for country-year pairs featuring these two years, zero otherwise). To address missing values (which occurred for all variables other than YEAR, INTERNET, COMPUTERS, MAINTEL, P1564, P65PLUS, URBAN and RISK which were fully populated), we used a standard imputation method: estimation of missing values via the regression of a predictor on the other predictors; a percentage of between 4.6% (for the variable TRADE) and 36.9% (for the variable ETHNIC) of the observations were imputed in this manner. After imputation, our usable sample size was 480, corresponding to 160 nations.

To attain an initial understanding of the relationship between the number of Internet users per 10,000 individuals and our predictors, we propose the following descriptive graphs in Figures 2a-f.

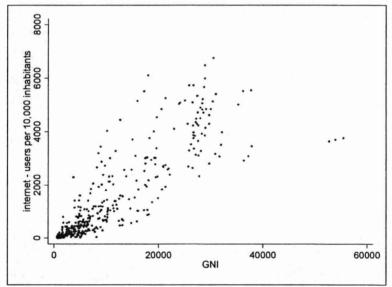


Figure 2a. Internet Users and GNI

The relationship (Figure 2a) between the number of Internet users and GNI shows some signs of non-linearity in the lower ranges of GNI, and some heteroskedasticity (uneven variances across the graph, resulting in a funnel shape; the number of Internet

users is higher – but also more variable – as GNI increases). Figure 2b reveals that GNI seems to have very little effect on the number of Internet users when schooling is very low, and that, more generally, the relationship between the number of Internet users and GNI seems to depend on schooling.

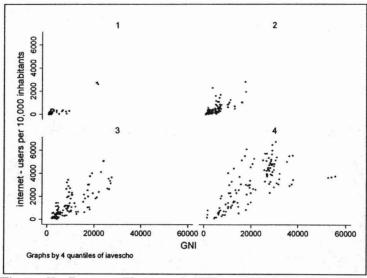


Figure 2b: Internet Users and GNI by SCHOOL Quaritiles

In Figure 2c, we can see that the curvature in the relationship between the number of Internet users and GNI is greater for lower quartiles of main telephone lines than for the top quartile of that variable.

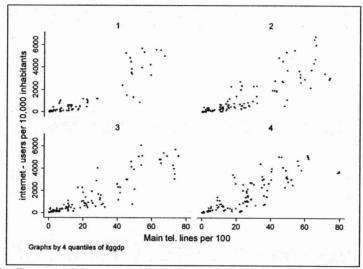


Figure 2c. Internet Users and Main Telephone Lines by TRADE Qaurtiles

The non-linear relationship between the number of Internet users and schooling emerges very clearly from Figure 2d.

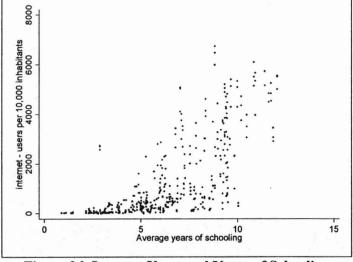
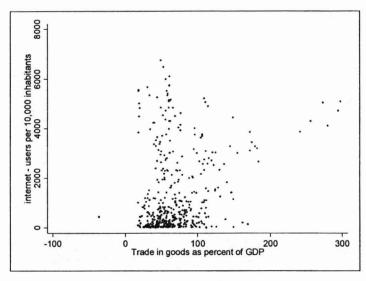


Figure 2d. Internet Users and Years of Schooling

Figure 2e displays the complicated relationship between the number of Internet users and the trade variable.



Figures 2e. Internet Users and Trade In Goods as Percent of GDP

Interestingly, as illustrated in Figure 2f, risk seems to have very little effect on the number of Internet users for lower quartiles of the main telephones variable, but a more pronounced effect in the upper quartiles.

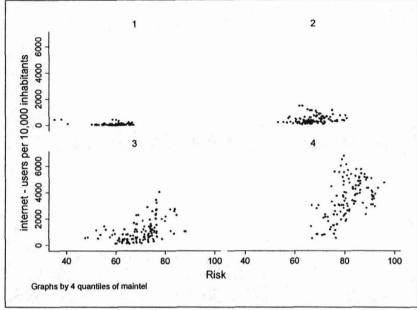


Figure 2f. Internet Users and Risk by MAINTEL Quartiles

In exploring the impact of determinants of the IDD, we embrace the complexity of each observation in space and time and attempt to grasp and explain that complexity using MARS.

This methodology allows us to identify whether the effect of some of our predictors on the number of Internet users might be non-linear; for instance, a glimpse at Figure 2d seems to indicate that the average number of years of schooling for adults does not have much effect on the number of Internet users until about 8 or 9 years of schooling at which point the effect becomes very strong. If one ignores the particular shape of this dependence, one might try to fit a straight line to this graph with inaccurate results. MARS therefore helps us identify break-points in the relationship between the dependent variable and our predictors which might not be obviously identifiable from an exploration of the data.

In addition, the methodology allows us to discover if the effect of one predictor on the dependent variable might depend on the level of another predictor. For example a quick glance at Figure 2f would seem to imply that the effect of RISK on INTERNET is essentially non-existent for low level of MAINTEL, but quite pronounced for high levels of MAINTEL. Our models help us identify such interactions effects even when they are not easy to detect with an exploration of the data.

The main ideas behind the algorithm are provided to the interested reader in the box below:

Main ideas of the MARS algorithm

The MARS algorithm, proposed by Friedmann in 1990, entails the following steps. For each continuous independent variable, MARS creates a piecewise linear function with too many change points (knots) to begin with, and then prunes unnecessary knots by a backward procedure.

- For each independent categorical variable, MARS merges categories if necessary to obtain a better fit with the dependent variable.
- MARS seeks suitable interactions between independent variables.
- MARS produces a series of transformations of independent variables that take into account non-linearities and interactions, known as Basis Functions.
- MARS then uses its Basis Functions as independent variables to estimate a least-squares model.

The MARS methodology essentially works by discovering judicious transformations of the independent variables (the Basis Functions) which take into account non-linear dependencies between the dependent variable and the predictors as well as interactions among the predictors. A list of Basis Functions (BFs) and the resulting least-squares equation is given in the Appendix.

Examining the definition of BF3 and BF16, one can see that two breaks occur in the variable on schooling (SCHOOL), one at about 10 years, and one at about 11 years. This implies that the slope of the variable SCHOOL will change in the model when SCHOOL passes through these values. Also, as evidenced for example in BF2, SCHOOL interacts with GNI, so that the effect of SCHOOL on the dependent variable will depend on GNI, and vice-versa.

Over-fitting is controlled in a number of ways in the MARS algorithm, which we do not detail here for the sake of space (details are given in Friedman, 1991). In our model, we used ten-fold cross-validation to determine the penalty to be attributed to each knot (break-point). We also note that our independent variables are correlated, but not to the extent of causing severe numerical or interpretation problems. For a discussion of multi-collinearity issues in the context of MARS models we refer the reader to the article by De Veaux and Ungar (1994). We finally note that we did not attempt to model the time autocorrelation between values of variables for 2001 and 2002-2003, because two time periods are likely to be too few to model this autocorrelation. Time effects and their interactions with suitable predictors (MAINTEL) are, however, captured by Basis Functions BF18, BF 19 and BF 21.

ANALYSIS AND INTERPRETATION

The four most important variables identified by the MARS algorithm, aside from YEAR (with a relative importance level of about 41) are as follows:

- income per capita ("GNI"), with a relative importance level of 100
- infrastructure ("MAINTEL"), with an importance level of about 73

- education ("SCHOOL"), with an importance level of about 54
- trade ("TRADE"), with an importance level of about 48.

Other variables that intervene in the model are P65PLUS (importance level of about 29), RISK (importance level of about 27) and URBAN (importance level of about 6.5). We note that the importance of a variable is measured by the amount of accuracy lost to the model if that variable is dropped. Our model, using 480 observations, yields an R-square of .93.

MARS output can be fairly complicated to interpret, because of the difficulty in interpreting interactions, particularly in the presence of non-linear relationships. Therefore, we propose the following graphical method. In order to explore the dependence of the number of Internet users on each of the four leading indicators, we fix the levels of the remaining three indicators at their first, second (median), and third quartile, respectively, and other variables at their median, providing a basis for comparison. In Figures 3-6, these levels are referred to as "111", "222", and "333", respectively. For example, in Figure 3b, the thick solid line refers to the dependence of Internet on GNI for levels of infrastructure, education, and trade ("IET") fixed at their third quartiles in 2002-03 (int33302).

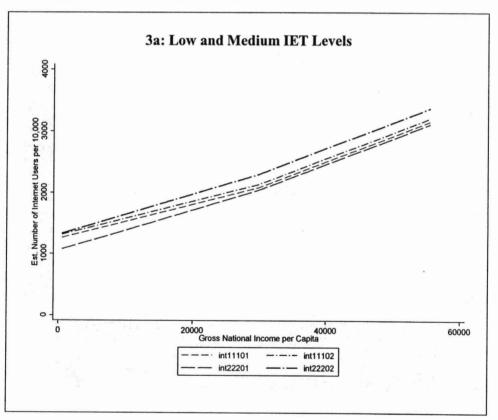
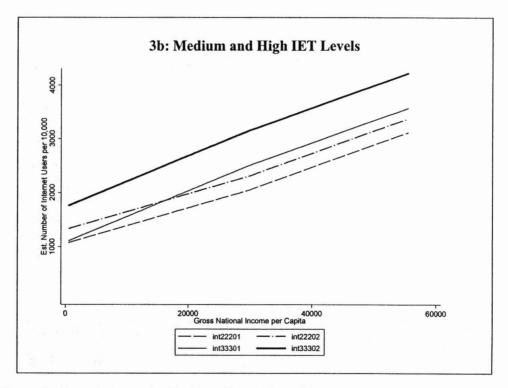


Figure 3. Effect of GNI on Internet Users, for Different Levels of IET (infrastructure, education and trade)

As shown in Figure 3a, the effect of GNI is more pronounced after \$30,000, and the break is most evident for the lower quartiles of our other key variables ("IET"). After \$30,000 the effect of GNI is similar for all levels.

Figure 3b shows that the difference between 2001 and 2002-03 is higher for higher quartiles of IET (solid lines) than for median levels (dashed lines): during this period, the growth in Internet users is highest in countries with the best infrastructure, highest education, and highest trade. Also, the nature of the break is different for third IET quartile: for this group, the effect of GNI is weaker after \$30,000.

In Figures 4a and 4b, the groups are determined by the quartiles of wealth, education, and trade ("WET"). Dependence of the number of Internet users on infrastructure at all levels is somewhat similar, but there is a slight break at about 62 and very sharp increase (clear break in Figure 4a and 4b) after 65 lines per 100 (close to the 95^{th} percentile of the MAINTEL variable). The models suggest that the number of telephone lines is a meaningful predictor for the number of Internet users only after the number of telephone lines has reached a high level, in particular for lower WET levels: for higher WET levels, the estimated number of Internet users starts increasing with infrastructure right from the start – albeit much more slowly than after the break point of 65 lines per 100 population.



Note: the lines designated with the suffix "01" are for 2001, and those designated with "02" are for 2002-03

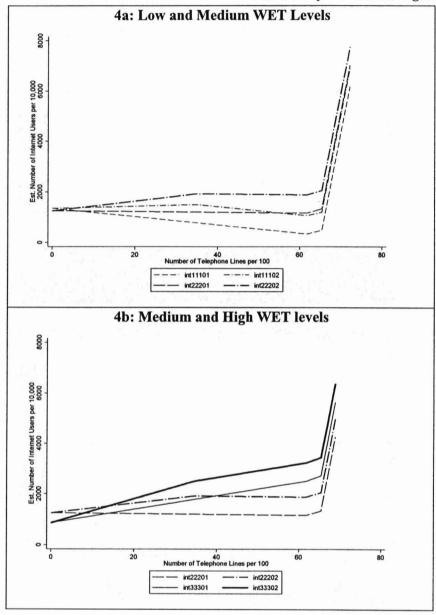


Figure 4. Effect of Telephone Lines on Internet Users, for Different Levels of WET (wealth, education and trade).

Note: Excluded from the graphs are any observations with estimated number of Internet users – 8000 (11 observations for a., 13 observations for b).

We also note that the slope of the lines after the break-point of 65 lines per 100 population actually depend on the value of the variable RISK. When RISK is fixed at its median value of 68.15, that slope is about 836, as represented in Figure 4. However, when RISK is fixed at its first quartile (61.91), the slope is as high as 1,122.6, and when RISK is fixed at its third quartile (76.33), the slope is quite a bit

lower at about 460. Interestingly this implies that a lower risk rating (we recall that less risky countries have higher values of this variable) is associated with a steeper increase in Internet with infrastructure, at least after a break-point of 65 lines per 100 population has been reached.

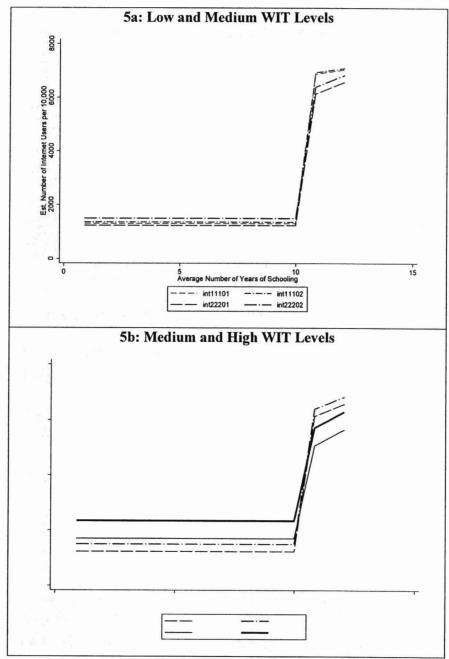


Figure 5. Effect of Schooling on Internet Users, for different levels of WIT (wealth, infrastructure and trade)

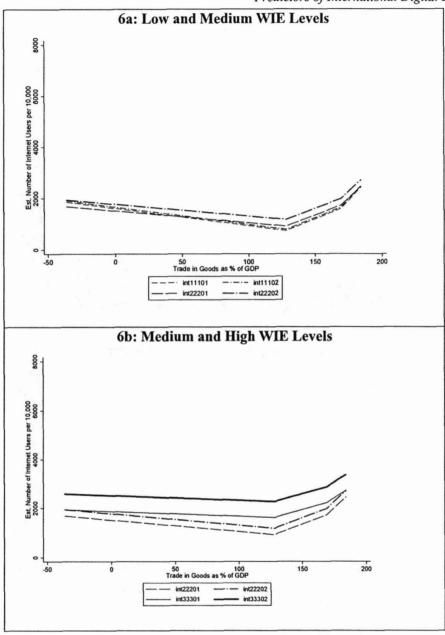
In Figures 5a and 5b, the groups are determined by wealth, infrastructure, and trade ("WIT"). Education appears to have no effect on Internet until about ten (9.99) average years of schooling (close to 93^{rd} percentile of the SCHOOL variable), at which point a very sharp increase takes place. At 10.8 years (96^{th} percentile of the SCHOOL variable), this effect tapers off. Beyond 10.8, the impact of education is stronger for lower levels of WIT, since the graphs for different levels of WIT in the high range of education are higher and higher as levels of WIT decrease. As in Figure 5a, the lower WIT quartiles (111) show a sharper initial post-ten year increase but also taper off faster than the higher WIT values in Figure 5b (222), which in turn behave the same way relative to the highest WIT values (333). As in the graphs of infrastructure (Figure 4) the difference between 01 and 02/03 is larger for the higher WITs (total range in Figure 5b) than the lower WITs (total range in Figure 5a).

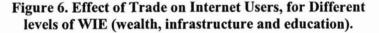
In Figures 6a and 6b, the groups are based on wealth, infrastructure, and education ("WIE"). The variable examined is trade in goods as a percentage of GDP, which represents a measure of a country's reliance on external trade or the openness of its economy. Using this measure, Internet use actually declines with trade until the point at which trade represents 128% of GDP (near the 94th percentile of TRADE). It then increases with trade as a percentage of GDP, accelerating dramatically after 169.05% (at the tail end of the TRADE distribution, past the 95th percentile).

An inversion of WIE (wealth, infrastructure, education) levels is evident in the relationship between trade and Internet use, as shown by both Figure 6a and 6b. In other words, the importance of trade at lower levels of WIE actually increases more dramatically than at higher levels. The greatest inversion is in Figure 6b, which means that at the highest levels of WIEs, trade has the weakest impact on Internet use.

DISCUSSION AND CONCLUSIONS

On the surface, it appears that our findings simply confirm what earlier research has discovered: the most important factors affecting the number of Internet users within a country are economic wealth and telecommunications infrastructure, followed by the level of education and the trade openness of the country. Our analytical approach allowed us, however, to discover associations that earlier research had missed. In particular, our results reveal that the effects of the main variables vary according to the levels of the other variables (e.g., the effects of wealth depend upon the aggregated levels of infrastructure, education, and trade). Also, the analytical approach we used allowed us to demonstrate that the effects of the main variables are seldom linear. The rest of this section will discuss the specific findings and their practical implications at a more detailed level.





As indicated by earlier research, the economic wellbeing of a nation (in our case measured by GNI) had a direct positive association with the proportion of Internet user in the population. This finding was not particularly surprising, as income measures represented the most frequently cited determinant in the comprehensive review of the literature by Dewan and Riggins (2005). A more detailed exploration of

the results suggests, however, that the income effects are not always the same: at a higher aggregated level of infrastructure, education, and trade openness (IET), the proportion of Internet users is higher than at the lower IET levels but the proportion also increased by a larger amount over the 2001-2003 period. We explored the change over only a three-year period and therefore, we cannot make inferences regarding longer-term trends but the short-term effect is clear. At the least advanced IET levels Internet use was barely advancing at all, regardless of the income level. Therefore, special policy measures to encourage growth in infrastructure, education, and trade may be needed to enable Internet usage growth in these countries.

Having a widespread traditional telecommunications infrastructure is a necessary prerequisite for Internet access: dial-up access is entirely based on the voice infrastructure, and a major portion of the broadband connections (those using Digital Subscriber Line or DSL technologies) are using the same underlying wiring as voice does. Thus, it is not surprising that the level of the telecommunications infrastructure is closely associated with the proportion of Internet users in the population. Our results suggest, however, that the effect is not linear: the infrastructure development process must reach a relatively high level before infrastructure improvements will be associated with increased levels of Internet use. In other words, if a critical level has not yet been reached, it is unlikely that small investments will lead to appreciable results. This suggests that political decision-makers and those responsible for infrastructure development have to be patient and not expect immediate advances in Internet use, if the starting level of infrastructure is low. It is also worth noting that at higher GNI, education, and trade (WET) levels infrastructure effects start earlier than at the lower WET levels, where the infrastructure has virtually no effect until a relatively high critical threshold is reached.

In regard to education, policy changes should also be implemented with patience. Higher levels of Internet use are associated with higher levels of education, but only after reaching a threshold of ten years of schooling. Moreover, increasing the level of education beyond this threshold level will not have a strong effect on Internet use: after about 11 years of education, the acceleration of Internet use drops off. Thus, it appears that the critical investment is to bring societies to the completion of the secondary level of education.

Our results have revealed that the relationship between openness to trade and Internet penetration is quite complicated. We were able to verify the results of Chinn and Fairlie (2004), who found that this variable has a negative effect on Internet penetration, a result they considered to be somewhat counter-intuitive. We also observed such a negative effect, but this effect is inverted once high levels of trade openness are reached. We note here that the variable on trade represents to some extent the size of the economy; very large values (in the 240 -280% range) occur for small economies such as Singapore and Hong Kong. By contrast, values for large economies such as the US or Canada, are approximately 18% and 66%, respectively.

It is essential to remember that the macro-level indicators do not necessarily reveal all factors that may potentially have an impact on Internet penetration in a specific country. For example, the effects of advances in technology are not limited to infrastructure technologies (such as the basic phone network that our study and many earlier ones have used). The availability of affordable access technologies, such as desktops, laptops and "smart phones", as well as other individual level factors identified by various authors and summarized by Dewan and Riggins (2005) also matter. This seams to have already been recognized by some, such as MIT's Nicolas Negroponte, who recently launched the "One Laptop Per Child" project (<u>http://laptop.org/</u>). One of the objectives of this project is to enhance Internet connectivity where the basic infrastructure is already present by growing the diffusion of affordable access devices. Therefore, it is important that future studies of the International Digital Divide consider the potential effects of affordable access technologies along with the macro-level determinants discussed in this paper and in the literature.

The main contribution of this study is an explicit recognition of the complexity of the relationships between Internet use and its determinants. Our results, obtained using the MARS methodology, clearly demonstrate that these relationships are not always linear and that the independent variables have significant interactions. While our results support the known importance of wealth, infrastructure, and education on Internet use, they take our understanding of these phenomena further by revealing at a more detailed level how these effects vary depending on the values of other variables. Moreover, our findings suggest that policy makers interested in increasing the Internet use, and thus closing the digital divide, should consider multiple factors simultaneously and understand their joint effects. They should also be mindful that determinants of Internet use must reach a critical level before their effect is realized.

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APPENDIX: MARS BASIS FUNCTIONS

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Basis Functions
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BF1 = max(0, GNI - 29740.002);
BF2 = max(0, 29740.002 - GNI);
BF3 = max(0, SCHOOL - 9.990) * BF2;
BF5 = max(0, MAINTEL - 65.370);
BF6 = max(0, 65.370 - MAINTEL );
BF7 = max(0, TRADE - 128.097) * BF2;
BF8 = max(0, 128.097 - TRADE) * BF2;
BF10 = max(0, 61.770 - MAINTEL) * BF2;
BF11 = max(0, P65PLUS - 15.545) * BF2;
BF14 = max(0, 169.050 - TRADE) * BF6;
BF15 = max(0, RISK - 35.025) * BF5;
BF16 = max(0, SCHOOL - 10.840) * BF6;
BF18 = (YEAR = 2001);
BF19 = (YEAR = 2002 \text{ OR } YEAR = 2003);
BF21 = max(0, 34.860 - MAINTEL) * BF19;
BF24 = max(0, P65PLUS - 15.172) * BF2;
BF30 = max(0, URBAN - 7.395) * BF1;
```

Y = 5247.620 - 0.062 * BF2 + 0.237 * BF3 + 2357.414 * BF5 - 107.878 * BF6 + 0.002 * BF7 - 0.001 * BF8 + 0.002 * BF10 - 0.175 * BF11 + 0.551 * BF14 - 45.929 * BF15 - 102.696 * BF16 - 719.146 * BF18 - 20.782 * BF21 + 0.105 * BF24 + .824533E-03 * BF30;

model INTERNET = BF2 BF3 BF5 BF6 BF7 BF8 BF10 BF11 BF14 BF15 BF16 BF18 BF21

BF24 BF30;

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